# Analytic Pricing of Volatility-Equity Options within Wishart-Based Stochastic Volatility Models 

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#### Abstract

We price for different affine stochastic volatility models some derivatives that recently appeared in the market. These products are characterized by payoffs depending on both stock and its volatility. We provide closed-form solution for different products and two multivariate Wishartbased stochastic volatility models. The methodology turns out to be independent of the dimension of the problem. Overall, our results highlight the great flexibility and tractability of Wishart-based stochastic volatility models to develop multivariate extensions of the Heston model.


Keywords: Option Pricing, Target Volatility Options, Corridor Variance Swap, Double Digital Call, Wishart Stochastic Volatility Models.

[^0]
## 1 Introduction

The first generation of equity derivative products had payoffs depending on the stock price, like vanilla options, or the stock price path, like look back or barrier options. During the nineties volatility has become an asset class by itself, first by the creation of the volatility index VIX and almost ten years later (around 2004) by the emergence and steady growth of VIX futures and VIX option markets. More recently, on a large number of indexes, corresponding volatility indexes were built and serve as underlying for derivatives, like futures and options. These evolutions have led to the development of exotic volatility derivatives, whose payoffs depend on the volatility path ${ }^{1}$, or equity-volatility derivatives, whose payoffs depend explicitly on both the stock and the volatility (path). The growing complexity of the equity-volatility derivative market has created new modelling and implementation challenges.

Among the recent equity-volatility products that have attracted some attention are the target volatility option (TVO), the corridor variance swap (CVS) and the double digital call option (DDC). Within the Heston (1993)'s model closed-form solutions were proposed by several authors ${ }^{2}$. A TVO is a European-type derivative contract whose value at maturity is given by the product of three terms: a vanilla European call, a target volatility parameter representing the investors expectation of the future realized volatility and the inverse of the realized volatility of the underlying. For this products, Di Graziano and Torricelli (2012) and Torricelli (2013) provide within the Heston (1993) framework a pricing methodology based on the Laplace transform, see also recent developments in Torricelli (2014). Grasselli and Marabel Romo (2015) considered in the 2 -factor Heston model the pricing of vanilla and forward-starting TVO, that is, TVO where the strike is determined at a later date. A Corridor Variance Swap is a generalisation of a standard variance swap in that the volatility is accumulated only when the underlying stock is within a pre-specified band, see Carr and Lewis (2004). In Zheng and Kwok (2014) the pricing of discrete corridor variance swap is investigated within a jump-diffusion Heston model by using the Fourier transform approach. Lastly, in Torricelli (2013) the DDC is considered within the Heston model and a closed form solution is proposed. All these results crucially exploit the analytical tractability of the Heston model and more generally of the standard affine framework. These products illustrate the growing importance of equity-volatility derivatives and underline the

[^1]need to understand whether they can be priced within more sophisticated frameworks that were recently developed.

The seminal work of Heston was extended to a multivariate stochastic volatility model using the vector affine process in Duffie et al. (2000) (see also Duffie and Kan (1996)) or by using two square root processes as in Christoffersen et al. (2009). Following the introduction in finance of the Wishart process, which is a matrix stochastic process, by Gouriéroux and Sufana (2010) more profound multivariate extensions of the affine model were proposed in Da Fonseca, Grasselli, and Tebaldi (2008) and Da Fonseca, Grasselli, and Tebaldi (2007). The first one is a multivariate stochastic volatility singlestock model while the second one is a multi-asset stochastic volatility and correlation model. These two models allow the computation in closed form of the characteristic function so that efficient option pricing through fast Fourier transform algorithms can be performed. However, extending results available for the Heston model to those more sophisticated models is far from being a straightforward task. Depending on the product at hand it may or may not be possible to price in closed form these products within Wishart-based stochastic volatility models. It is therefore of interest to understand when such extensions can be performed.

In this work we propose a general pricing framework for volatility derivatives based on a simple yet powerful approach which combines conditioning with respect to the subfiltration generated by the volatility path and Fourier techniques. This conditioning technique is standard in option pricing, see Leblanc (1996) or Henry-Labordère (2009), but our work will underline its importance for handling multi-factor or multi-asset stochastic volatility models. We provide closed-form solutions for the TVO price based on the Fourier transform much in the spirit of Torricelli (2013). For the corridor variance swap we develop a pricing formula in the spirit of Zheng and Kwok (2014) and for the Double Digital call option we show how a closed-form solution can be obtained. The essential contribution of our work is to explain how these techniques apply to Wishart based stochastic volatility models, either the WASC of Da Fonseca et al. (2007) or the WMSV of Da Fonseca et al. (2008). In these particular cases the closed form solution for the characteristic function turns out to be crucial for an efficient numerical implementation. Within the general affine framework we price these three products for typical parameter values (these values are obtained from a vanilla option calibration procedure). This will allow us to illustrate an important problem of exotic option pricing, namely the issue of model risk, and provide an integrated perspective of exotic option pricing and calibration on vanilla options with important consequences in terms of regulation of derivative products.

The structure of the paper is as follows. In Section 1 we present the different models; in Section 2 we focus on the pricing of TVO and corridor variance swap; in Section 3 we provides a numerical implementation; Section 4 contains the pricing of a digital call; Section 5 provide some open problems illustrating some intrinsic difficulties related to Wishart based models. The last section concludes and we gather all tables in the Appendix.

## 2 The Models

In this section we briefly review the stochastic volatility models that will be considered in the sequel together with their moment generating functions. We present the Heston, the BiHeston, the WMSV and the WASC models. The first two are well known but are given here for convenience as they will be involved in the numerical experiments. We could have unified the presentation of the Heston and BiHeston models but we prefer to avoid cumbersome notations.

### 2.1 The Heston (1993) Model

We denote by $s_{t}$ a stock whose dynamics are given by the following system of stochastic differential equations (SDEs in the sequel):

$$
\begin{align*}
& d s_{t}=s_{t} r d t+s_{t} \sqrt{v_{t}}\left(\rho d w_{1, t}+\sqrt{1-\rho^{2}} d w_{2, t}\right), s_{0}>0,  \tag{1}\\
& d v_{t}=\kappa\left(\theta-v_{t}\right) d t+\sigma \sqrt{v_{t}} d w_{1, t}, v_{0}>0, \tag{2}
\end{align*}
$$

where $w_{t}=\left(w_{1, t}, w_{2, t}\right)_{t \geq 0}$ is a two-dimensional Brownian motion, $\kappa \in \mathbb{R}, \kappa \theta \in \mathbb{R}_{+}, \sigma>0$ and $\rho \in[-1,1]$.

The joint moment-generating function, defined by $G_{\text {HES }}\left(t, z, \lambda_{v}, \Lambda_{v}\right)=\mathbb{E}\left[e^{z \ln s_{t}+\lambda_{v} v_{t}+\Lambda_{v} \int_{0}^{t} v_{u} d u}\right]^{3}$, is known in closed form. In fact, the affine property of the model leads to the following lemma whose standard proof is omitted.

Lemma 2.1. The moment generating function of $\left(\ln s_{t}, v_{t}, \int_{0}^{t} v_{u} d u\right)$ is given by:

$$
\begin{aligned}
G_{\mathrm{HES}}\left(t, z, \lambda_{v}, \Lambda_{v}\right) & =\mathbb{E}\left[e^{z \ln s_{t}+\lambda_{v} v_{t}+\Lambda_{v}} \int_{0}^{t} v_{u} d u\right. \\
& =e^{z \ln s_{0}+z r t+A(t) v_{0}+b(t)},
\end{aligned}
$$

[^2]with the deterministic functions $A(t), b(t)$ defined as:
\[

$$
\begin{aligned}
A(t) & =\frac{\eta \lambda_{+} e^{-\sqrt{\Gamma} t}+\lambda_{-}}{\frac{\sigma^{2}}{2}\left(\eta e^{-\sqrt{\Gamma} t}+1\right)} \\
b(t) & =\frac{2 \kappa \theta}{\sigma^{2}}\left(t \lambda_{-}-\log \left(\frac{\eta e^{-\sqrt{\Gamma} t}+1}{1+\eta}\right)\right)
\end{aligned}
$$
\]

with

$$
\begin{align*}
\lambda_{ \pm} & =\frac{(\kappa-z \rho \sigma) \pm \sqrt{\Gamma}}{2}  \tag{3}\\
\Gamma & =(\kappa-z \rho \sigma)^{2}-\sigma^{2}\left(z^{2}-z+2 \Lambda_{v}\right)  \tag{4}\\
\eta & =-\frac{\sigma^{2} \lambda_{v}-2 \lambda_{-}}{\sigma^{2} \lambda_{v}-2 \lambda_{+}} \tag{5}
\end{align*}
$$

In the Heston model the stock's variance, the variance of stock's variance and the correlation between the log-stock and its (instantaneous) variance are given by

$$
\begin{align*}
d\left\langle\ln s_{t}\right\rangle & =v_{t} d t  \tag{6}\\
d\left\langle\operatorname{vAR}\left(s_{t}\right)\right\rangle & =d\left\langle v_{t}\right\rangle=\sigma^{2} v_{t} d t  \tag{7}\\
d \operatorname{CoRR}\left(\ln s_{t}, \operatorname{vAR}\left(s_{t}\right)\right) & =\rho d t \tag{8}
\end{align*}
$$

These results are all well known but they will allow us to underline the specificities and contributions of the next models.

### 2.2 The BiHeston Model

We consider here the Christoffersen et al. (2009) specification of a model where the diffusion term of the asset is described as a combination of two square root processes. This specification is also referred to as the Double Heston, or BiHeston model. The stock price dynamics are defined via the following set of stochastic differential equations:

$$
\begin{align*}
& d s_{t}=s_{t} r d t+s_{t}\left(\sqrt{v_{t}^{0}} d Z_{t}^{0}+\sqrt{v_{t}^{1}} d Z_{t}^{1}\right), s_{0}>0  \tag{9}\\
& d v_{t}^{0}=\kappa_{0}\left(\theta_{0}-v_{t}^{0}\right) d t+\sigma_{0} \sqrt{v_{t}^{0}} d W_{t}^{0}, v_{0}^{0}>0  \tag{10}\\
& d v_{t}^{1}=\kappa_{1}\left(\theta_{1}-v_{t}^{1}\right) d t+\sigma_{1} \sqrt{v_{t}^{1}} d W_{t}^{1}, v_{0}^{1}>0 \tag{11}
\end{align*}
$$

with $d\left\langle Z^{0}, W^{0}\right\rangle_{t}=\rho_{0} d t, d\left\langle Z^{1}, W^{1}\right\rangle_{t}=\rho_{1} d t$, while all other correlations are set to zero in order to grant the analytical tractability of the model ${ }^{4}$. The parameters in (10) and (11) satisfy the following

[^3]restrictions: $\kappa_{i} \in \mathbb{R}, \kappa_{i} \theta_{i} \in \mathbb{R}_{+}, \sigma_{i}>0$ and $\rho_{i} \in[-1,1]$ for $i=\{0,1\}$.

The joint moment generating function of the asset returns, the variance process $v_{t}=\left(v_{t}^{0}+v_{t}^{1}\right)$ and the integrated variance process $V_{t}=\int_{0}^{t} v_{u} d u=\int_{0}^{t}\left(v_{u}^{0}+v_{u}^{1}\right) d u$ is given by:

$$
G_{2 \mathrm{HES}}\left(t, z, \lambda_{v^{0}}, \lambda_{v^{1}}, \Lambda_{v^{0}}, \Lambda_{v^{1}}\right)=\mathbb{E}\left[e^{z \ln s_{t}+\lambda_{v^{0}} v_{t}^{0}+\lambda_{v^{1}} v_{t}^{1}+\int_{0}^{t}\left(\Lambda_{v^{0}} v_{u}^{0}+\Lambda_{v^{1}} v_{u}^{1}\right) d u}\right]
$$

Since the model is affine, it is natural to look for an exponentially affine form, and the next lemma gives the explicit expression for this function:

Lemma 2.2. The joint moment generating function of $\left(\ln s_{t}, v_{t}^{0}, v_{t}^{1}, \int_{0}^{t} v_{u}^{0} d u, \int_{0}^{t} v_{u}^{1} d u\right)$ is given by:

$$
\begin{align*}
G_{2 \mathrm{HES}}\left(t, z, \lambda_{v^{0}}, \lambda_{v^{1}}, \Lambda_{v^{0}}, \Lambda_{v^{1}}\right) & =\mathbb{E}\left[e^{z \ln s_{t}+\lambda_{v 0} v_{t}^{0}+\lambda_{v^{1}} v_{t}^{1}+\int_{0}^{t}\left(\Lambda_{v 0} v_{u}^{0}+\Lambda_{v^{1}} v_{u}^{1}\right) d u}\right] \\
& =e^{z x_{0}+z r t+A^{0}(t) v_{0}^{0}+b^{0}(t)+A^{1}(t) v_{0}^{1}+b^{1}(t)} \tag{12}
\end{align*}
$$

where the deterministic functions $A^{j}, b^{j}, j=0,1$, satisfy:

$$
\begin{aligned}
A^{j}(t) & =\frac{\eta_{j} \lambda_{+}^{j} e^{-\sqrt{\Gamma_{j}} t}+\lambda_{-}^{j}}{\frac{\sigma_{j}^{2}}{2}\left(\eta_{j} e^{-\sqrt{\Gamma_{j}} t}+1\right)} \\
b^{j}(t) & =\frac{2 \kappa_{j} \theta_{j}}{\sigma_{j}^{2}}\left(t \lambda_{-}^{j}-\log \left(\frac{\eta_{j} e^{-\sqrt{\Gamma_{j}} t}+1}{1+\eta_{j}}\right)\right)
\end{aligned}
$$

with

$$
\begin{align*}
\lambda_{ \pm}^{j} & =\frac{\left(\kappa_{j}-\rho_{j} \sigma_{j} z\right) \pm \sqrt{\Gamma_{j}}}{2}  \tag{13}\\
\Gamma_{j} & =\left(\kappa_{j}-\rho_{j} \sigma_{j} z\right)^{2}-\sigma_{j}^{2}\left(z^{2}-z+2 \Lambda_{v^{j}}\right)  \tag{14}\\
\eta_{j} & =-\frac{\sigma_{j}^{2} \lambda_{v^{j}}-2 \lambda_{-}^{j}}{\sigma_{j}^{2} \lambda_{v^{j}}-2 \lambda_{+}^{j}} \tag{15}
\end{align*}
$$

This model constitutes a multivariate extension of the Heston model and uses two unrelated square root processes. It would be possible to use instead the vector affine process of Duffie et al. (2000) (see also Duffie and Kan (1996)). However, in that case the moment-generating function would involve Riccati ordinary differential equations that can not be computed in closed form (see Grasselli and Tebaldi (2008) for further details regarding the solvability of these equations) and require the use of numerical schemes implying a much higher computational complexity. Furthermore, if the derivative with respect to a model parameter or the argument of the moment-generating function is needed then the computational burden is even higher. As a consequence, multidimensional extensions of Heston's
model can come with a significant (complete) loss of analytical tractability. The fact that the two square root processes cannot be correlated is an intrinsic constraint of the Duffie-Kan affine processes and one of the main advantages of the next model is to remove that constraint.

For this model the stock's variance, the variance of stock's variance and the correlation between the log-stock and its (instantaneous) variance are given by

$$
\begin{align*}
d\left\langle\ln s_{t}\right\rangle & =v_{t}^{0}+v_{t}^{1} d t,  \tag{16}\\
d\left\langle\operatorname{vAR}\left(s_{t}\right)\right\rangle & =\sigma_{0}^{2} v_{t}^{0}+\sigma_{1}^{2} v_{t}^{1} d t,  \tag{17}\\
d \operatorname{CoRR}\left(\ln s_{t}, \operatorname{VAR}\left(s_{t}\right)\right) & =\frac{\rho_{0} \sigma_{0} v_{t}^{0}+\rho_{1} \sigma_{1} v_{t}^{1}}{\sqrt{v_{t}^{0}+v_{t}^{1}} \sqrt{\sigma_{0}^{2} v_{t}^{0}+\sigma_{1}^{2} v_{t}^{1}}} d t . \tag{18}
\end{align*}
$$

Equation (16) illustrates the fact that the BiHeston model is a multivariate extension of the Heston model and, in the present case, a two-factor extension. Compared to Equation (8), which displays a constant instantaneous correlation, Equation (18) underlines an interesting property of the BiHeston model. Namely, the spot-variance correlation is stochastic. Notice, however, that if $\rho_{0}$ and $\rho_{1}$ are negative, and that will happen in practice, then this correlation is of constant sign (negative).

### 2.3 The WMSV Model

We consider now the Wishart Multidimensional Stochastic Volatility Model (WMSV hereafter) proposed by Da Fonseca, Grasselli, and Tebaldi (2008). This stochastic volatility model extends the original Heston (1993) model to the case where the volatility is described by the Wishart process, a matrix-valued stochastic process introduced by Bru (1991). Within the WMSV model the dynamics for the stock price are given by the following SDE:

$$
\begin{equation*}
d s_{t}=s_{t} r d t+s_{t} \operatorname{Tr}\left[\sqrt{\Sigma_{t}}\left(d W_{t} R^{\top}+d B_{t} \sqrt{\mathbb{I}-R R^{\top}}\right)\right], s_{0}>0 \tag{19}
\end{equation*}
$$

where $\operatorname{Tr}$ is the trace operator, $W_{t}, B_{t} \in M_{n}$ (the set of square matrices) are composed by $n^{2}$ independent Brownian motions under the risk-neutral measure ( $B_{t}$ and $W_{t}$ are independent), $R \in M_{n}$ represents the correlation matrix and $\Sigma_{t}$ belongs to the set of symmetric $n \times n$ positive semi-definite matrices. In this specification the volatility is multi-dimensional and depends on the elements of the matrix process $\Sigma_{t}$, which is assumed to satisfy the following dynamics:

$$
\begin{equation*}
d \Sigma_{t}=\left(\Omega \Omega^{\top}+M \Sigma_{t}+\Sigma_{t} M^{\top}\right) d t+\sqrt{\Sigma_{t}} d W_{t} Q+Q^{\top}\left(d W_{t}\right)^{\top} \sqrt{\Sigma_{t}} \tag{20}
\end{equation*}
$$

with $\sqrt{ } \cdot$ denoting the matrix square root, initial condition $\Sigma_{0}$ a strictly positive definite matrix and parameters $\Omega, M \in M_{n}$, and $Q \in G L(n)$, the set of invertible $n \times n$ matrices.

Equation (20) characterizes the Wishart process investigated by Bru (1991) and then introduced in finance by Gouriéroux and Sufana (2010) and many other authors including Gouriéroux et al. (2009), Grasselli and Tebaldi (2008), Da Fonseca et al. (2008), Da Fonseca et al. (2007) ${ }^{5}$. For an extension of the classical work of Bru (1991), see for example Cuchiero et al. (2011). Existence and uniqueness results for the SDE (20) are provided in Mayerhofer et al. (2011). The Wishart processes represents the matrix analogue of the square root mean-reverting process. In order to grant the typical mean reverting feature of the volatility, the matrix $M$ is assumed to be negative semi-definite. The constant drift part satisfies $\Omega \Omega^{\top}=\beta Q^{\top} Q$ with the real parameter $\beta \geq n-1$ (see Cuchiero et al. (2011)) ${ }^{6}$. If $\beta$ satisfies the stronger assumption $\beta \geq n+1$ then the unique strong solution to the $\operatorname{SDE}$ (20) evolves as a strictly positive definite matrix, see Mayerhofer et al. (2011).

In this model the instantaneous variance of the asset returns is associated to the trace of the Wishart matrix, that is:

$$
\begin{equation*}
d\left\langle\ln s_{t}\right\rangle=\operatorname{Tr}\left[\Sigma_{t}\right] d t, \tag{21}
\end{equation*}
$$

which alone is not Markovian. Computing the expectation of this trace using a partial differential approach would also require consideration of the full state variable $\Sigma$.

For this model the stock's variance is given by Equation (21) and constitutes also a multivariate extension of the Heston model, while the variance of stock's variance and the correlation between the log-stock and its (instantaneous) variance are given by (in the particular case of $n=2$ )

$$
\begin{equation*}
d\left\langle\operatorname{vaR}\left(s_{t}\right)\right\rangle=Q_{11}^{2} \Sigma_{t}^{11}+Q_{22}^{2} \Sigma_{t}^{22} d t \tag{22}
\end{equation*}
$$

$$
\begin{align*}
d \operatorname{CorR}\left(\ln s_{t}, \operatorname{vaR}\left(s_{t}\right)\right) & =\frac{\operatorname{Tr}\left[R Q \Sigma_{t}\right]}{\sqrt{\operatorname{Tr}\left[\Sigma_{t}\right.} \sqrt{\operatorname{Tr}\left[Q^{\top} Q \Sigma_{t}\right]}} d t \\
& =\frac{R_{11} Q_{11} \Sigma_{t}^{11}+R_{22} Q_{22} \Sigma_{t}^{22}}{\sqrt{\Sigma_{t}^{11}+\Sigma_{t}^{22}} \sqrt{Q_{11}^{2} \Sigma_{t}^{11}+Q_{22}^{2} \Sigma_{t}^{22}}}+\frac{Q_{22} R_{12} \Sigma_{t}^{12}}{\sqrt{\Sigma_{t}^{11}+\Sigma_{t}^{22}} \sqrt{Q_{11}^{2} \Sigma_{t}^{11}+Q_{22}^{2} \Sigma_{t}^{22}}} d t \tag{23}
\end{align*}
$$

[^4]The correlation between the log-stock and its (instantaneous) variance, given by Equation (23), underlines a fundamental difference with the BiHeston model (the corresponding relation is Equation (18)). In fact, there is an additional factor $\Sigma_{t}^{12}$ that affects the correlation but not the stock's variance. Building an affine model using the standard affine framework with this property is a difficult task. Notice also that $\Sigma_{t}^{12}$ is stochastic and can be of any sign, thus relaxing the intrinsic constraint of the BiHeston model.

Given a scalar $z$ and two square (symmetric) matrices $\Lambda_{\Sigma}, \Lambda_{I}$, the joint moment generating function of the asset returns, the (Wishart) process $\Sigma_{t}$ and the integrated Wishart process $\int_{0}^{t} \Sigma_{u} d u$ is given by the function $G_{\mathrm{WMSV}}\left(t, z, \Lambda_{\Sigma}, \Lambda_{I}\right)=\mathbb{E}\left[e^{z \ln s_{t}+\operatorname{Tr}\left[\Lambda_{\Sigma} \Sigma_{t}\right]+\operatorname{Tr}\left[\Lambda_{I} \int_{0}^{t} \Sigma_{u} d u\right]}\right]$ which is known in closed form. Da Fonseca, Grasselli, and Tebaldi (2008) proved the following result:

Lemma 2.3. The joint moment-generating function of $\left(\ln s_{t}, \Sigma_{t}, \int_{0}^{t} \Sigma_{u} d u\right)$ is given by:

$$
\begin{equation*}
G_{\mathrm{WMSV}}\left(t, z, \Lambda_{\Sigma}, \Lambda_{I}\right)=e^{z \ln s_{0}+z r t+\operatorname{Tr}\left[A(t) \Sigma_{0}\right]+b(t)} \tag{24}
\end{equation*}
$$

where the deterministic matrix function $A(t)$ and the scalar function $b(t)$ satisfy the following $O D E$ (ordinary differential equations) ${ }^{7}$ :

$$
\begin{align*}
\frac{d A}{d t} & =A\left(M+z Q^{\top} R^{\top}\right)+\left(M+z Q^{\top} R^{\top}\right)^{\top} A+2 A Q^{\top} Q A+\frac{z(z-1)}{2} \mathbb{I}_{n}+\Lambda_{I},  \tag{25}\\
\frac{d b}{d t} & =\operatorname{Tr}\left[\Omega \Omega^{\top} A\right], \tag{26}
\end{align*}
$$

with boundary conditions $A(0)=\Lambda_{\Sigma}, b(0)=0$ and $\mathbb{I}_{n}$ is the identity matrix of $M_{n}$. The solution is explicitly given by:

$$
\begin{align*}
A(t) & =\left(\Lambda_{\Sigma} A_{12}(t)+A_{22}(t)\right)^{-1}\left(\Lambda_{\Sigma} A_{11}(t)+A_{21}(t)\right),  \tag{27}\\
b(t) & =-\frac{\beta}{2} \operatorname{Tr}\left[\log \left(\Lambda_{\Sigma} A_{12}(t)+A_{22}(t)\right)+t\left(M+z Q^{\top} R^{\top}\right)\right], \tag{28}
\end{align*}
$$

with

$$
\left(\begin{array}{ll}
A_{11}(t) & A_{12}(t)  \tag{29}\\
A_{21}(t) & A_{22}(t)
\end{array}\right)=\exp t\left(\begin{array}{ll}
M+z Q^{\top} R^{\top} & -2 Q^{\top} Q \\
\frac{z(z-1)}{2} \mathbb{I}_{n}+\Lambda_{I} & -\left(M+z Q^{\top} R^{\top}\right)^{\top}
\end{array}\right) .
$$

In order to compute some derivative prices we need to be able to differentiate the moment generating function. Thanks to the strong analytical tractability of the WMSV model this quantity can be computed explicitly as shown in the following result.

[^5]Corollary 2.1. The derivative of the function $g(\alpha):=G_{\mathrm{WMSV}}\left(t, z, \alpha \Lambda_{\Sigma}, \Lambda_{I}\right)$ (with $\alpha \in \mathbb{R}$ ) is given by

$$
\begin{equation*}
\frac{d g(\alpha)}{d \alpha}=\left(\operatorname{Tr}\left[\partial_{\alpha} A(t) \Sigma_{0}\right]+\partial_{\alpha} b(t)\right) G_{\mathrm{WMSV}}\left(t, z, \alpha \Lambda_{\Sigma}, \Lambda_{I}\right) \tag{30}
\end{equation*}
$$

with

$$
\begin{align*}
\partial_{\alpha} A(t) & =-\left(\alpha \Lambda_{\Sigma} A_{12}(t)+A_{22}(t)\right)^{-1} \Lambda_{\Sigma} A_{12}(t) A(t)+\left(\alpha \Lambda_{\Sigma} A_{12}(t)+A_{22}(t)\right)^{-1} \Lambda_{\Sigma} A_{11}(t)  \tag{31}\\
\partial_{\alpha} b(t) & =-\frac{\beta}{2} \operatorname{Tr}\left[\Lambda_{\Sigma} A_{12}(t)\left(\alpha \Lambda_{\Sigma} A_{12}(t)+A_{22}(t)\right)^{-1}\right] \tag{32}
\end{align*}
$$

Proof. Consider an invertible matrix $A$ of size $(n \times n)$ depending on the parameter $\alpha$. Then taking the derivative of $A A^{-1}=\mathbb{I}_{n}$ gives $\partial_{\alpha} A^{-1}=-A^{-1} \partial_{\alpha} A A^{-1}$ and using this equality from (27) we obtain (31). Let us now consider the map $\alpha \rightarrow \operatorname{Tr}\left[\log \left(\alpha \Lambda_{\Sigma} A_{12}(t)+A_{22}(t)\right)\right]$. Let $C=\alpha \Lambda_{\Sigma} A_{12}(t)+A_{22}(t)$ and $B=C-\mathbb{I}_{n}$ then we have

$$
\begin{aligned}
\operatorname{Tr}\left[\partial_{\alpha} \ln C\right] & =\operatorname{Tr}\left[\partial_{\alpha} \ln \left(\mathbb{I}_{n}+B\right)\right]=\operatorname{Tr}\left[\partial_{\alpha}\left\{B-\frac{B^{2}}{2}+\frac{B^{3}}{3}\right\}\right]=\operatorname{Tr}\left[\partial_{\alpha} B-\frac{\partial_{\alpha} B B+B \partial_{\alpha} B}{2}+\ldots\right] \\
& =\operatorname{Tr}\left[\partial_{\alpha} B-\partial_{\alpha} B B+\frac{\partial_{\alpha} B B^{2}}{2}+\ldots\right]=\operatorname{Tr}\left[\partial_{\alpha} B\left\{\mathbb{I}_{n}-B+\frac{B^{2}}{2}+\ldots\right\}\right]=\operatorname{Tr}\left[\partial_{\alpha} C C^{-1}\right]
\end{aligned}
$$

From this last equality we deduce the result.

This corollary illustrates the high tractability of the WMSV model. Equations (25) and (26) are useful as they underline the importance of $(27),(28)$ and (29). Had these last equations not been available, to compute the solution of Corollary 2.1 we would have had to discretize both the matrix ODE (25) and (26) as well as the sensitivity of these equations with respect to the parameters of interest. The computational cost would have been much higher.

### 2.4 The WASC Model

The Wishart Affine Stochastic Correlation (WASC hereafter) model of Da Fonseca, Grasselli, and Tebaldi (2007) consists in a $n$-dimensional risky asset $s_{t}=\left(s_{t}^{1}, . ., s_{t}^{n}\right)^{\top}$ whose dynamics are given by:

$$
\begin{equation*}
d s_{t}=\operatorname{diag}\left[s_{t}\right]\left[r \mathbf{1}+\sqrt{\Sigma_{t}} d Z_{t}\right] \tag{33}
\end{equation*}
$$

where $Z_{t} \in \mathbb{R}^{n}$ is a vector Brownian motion and 1 is a $n \times 1$ vector of ones, while the returns' variance-covariance matrix $\Sigma_{t}$ evolves stochastically, according to the Wishart dynamics (20) introduced previously.

The leverage effects and the asymmetric correlation effects are modeled by introducing the following correlation structure among Brownian motions:

$$
d Z_{t}=\sqrt{1-\rho^{\top} \rho} d B_{t}+d W_{t} \rho
$$

where $\rho$ is a vector of size $n$, with $\rho \in[-1,1]^{n}$ and $\rho^{\top} \rho \leq 1\left(B_{t}\right.$ is a vector Brownian motion under the risk-neutral measure and is independent of $W_{t}$ ). Remarkably, such correlation structure is the only one which is compatible with the affine property of the model, see Da Fonseca et al. (2007).

The instantaneous variance of the asset returns is associated to the diagonal terms of the Wishart matrix, that is:

$$
\begin{equation*}
d\left\langle\ln s_{t}^{i}\right\rangle=\Sigma_{t}^{i i} d t \tag{34}
\end{equation*}
$$

so that the integrated variance of the $i$-th asset is given by $V_{t}^{i}=\int_{0}^{t} \Sigma_{u}^{i i} d u$, thus leading to a structure very similar to the single factor Heston model case. It is important to notice, however, that even if the volatility structure for each asset is very simple compared to the BiHeston or WMSV models, the WASC model is a multiple-asset framework where assets are related in a non trivial way through the variance/covariance matrix. In fact, the instantaneous assets' covariance is given by $d\left\langle\ln s_{t}^{i}, \ln s_{t}^{j}\right\rangle=\Sigma_{t}^{i j} d t^{8}$. Notice that the (instantaneous) variance of the $i^{t h}$ asset when considered alone is not Markovian, so that a partial differential equation approach involving this state variable must also consider all other components of the matrix $\Sigma_{t}$. The WASC model can therefore be seen as a natural extension of the Heston model to a multiple-asset market where not only volatilities are stochastic, but also correlations among assets are random. What is more, the model is tractable due to the affine structure. To better understand the similarities and differences with the previous models, let us report, in addition to the stock's variance given by Equation (34), the variance of the stock's variance and the correlation between the log-stock and its variance for the particular case $n=2$ (we consider the first stock but similar equations apply for the second stock)

$$
\begin{align*}
d\left\langle\operatorname{vAR}\left(s_{t}^{1}\right)\right\rangle & =4\left(Q_{11}^{2}+Q_{21}^{2}\right) \Sigma_{t}^{11} d t  \tag{35}\\
d \operatorname{CoRR}\left(\ln s_{t}^{1}, \operatorname{vaR}\left(s_{t}^{1}\right)\right) & =\frac{\rho_{1} Q_{11}+\rho_{2} Q_{21}}{\sqrt{Q_{11}^{2}+Q_{21}^{2}}} d t \tag{36}
\end{align*}
$$

Comparing Equations (34), (35) and (36) with their Heston counterparts given by Equations (6), (7) and (8) clearly shows the similarities between the stock dynamics in the WASC and Heston

[^6]models. For example, it is possible to establish that $\Sigma_{t}^{11}$, for a given $t$, follows a noncentral chisquared distribution just like $v_{t}$ in the Heston model (see Da Fonseca and Grasselli (2011)), which provides a natural "mapping" strategy between the two models. But this similarity stands only for the marginal (instantaneous) variance distribution at a given time and is far from sufficient for having same stock price distributions. Let us stress again the fact that the process $\Sigma_{t}^{11}$ alone is not Markov. The dynamics of the first asset depends on parameters that are also involved in the dynamics of the second asset (if we consider two assets). As an example, the correlation between the second asset and its (instantaneous) variance is given by
\[

$$
\begin{equation*}
d \operatorname{CorR}\left(\ln s_{t}^{2}, \operatorname{vaR}\left(s_{t}^{2}\right)\right)=\frac{\rho_{1} Q_{12}+\rho_{2} Q_{22}}{\sqrt{Q_{12}^{2}+Q_{22}^{2}}} d t \tag{37}
\end{equation*}
$$

\]

from which we can see that the two asset dynamics share some parameters.

Given a vector $z \in \mathbb{R}^{n}$ and two square (symmetric) matrices $\Lambda_{\Sigma}, \Lambda_{I}$, the joint moment generating function of the asset returns, the (Wishart) variance process $\Sigma_{t}$ and the integrated variance process $\int_{0}^{t} \Sigma_{u} d u$ is given by the function $G_{\mathrm{WASC}}\left(t, z, \Lambda_{\Sigma}, \Lambda_{I}\right)=\mathbb{E}\left[e^{z^{\top} \ln s_{t}+\operatorname{Tr}\left[\Lambda_{\Sigma} \Sigma_{t}\right]+\operatorname{Tr}\left[\Lambda_{I} \int_{0}^{t} \Sigma_{u} d u\right]}\right]$ which is known in closed form (see Da Fonseca, Grasselli, and Tebaldi (2007) for the proof of the following result).

Lemma 2.4. The joint moment generating function of $\left(\ln s_{t}, \Sigma_{t}, \int_{0}^{t} \Sigma_{u} d u\right)$ is given by:

$$
\begin{equation*}
G_{\mathrm{WASC}}\left(t, z, \Lambda_{\Sigma}, \Lambda_{I}\right)=e^{z^{\top} \ln s_{0}+z^{\top} \mathbf{1} r t+\operatorname{Tr}\left[A(t) \Sigma_{0}\right]+b(t)} \tag{38}
\end{equation*}
$$

where the deterministic matrix function $A(t)$ and the scalar function $b(t)$ satisfy the following ODEs:

$$
\begin{align*}
\frac{d A}{d t} & =A\left(M+Q^{\top} \rho z^{\top}\right)+\left(M+Q^{\top} \rho z^{\top}\right)^{\top} A+2 A Q^{\top} Q A+\frac{1}{2} z z^{\top}-\frac{1}{2} \sum_{j=1}^{n} z^{j} e^{j j}+\Lambda_{I},  \tag{39}\\
\frac{d b}{d t} & =\operatorname{Tr}\left[\Omega \Omega^{\top} A\right], \tag{40}
\end{align*}
$$

with boundary conditions $A(0)=\Lambda_{\Sigma}, b(0)=0$ and $e^{j j}$ is the canonical basis of $M_{n}$. The solution is explicitly given by:

$$
\begin{align*}
A(t) & =\left(\Lambda_{\Sigma} A_{12}(t)+A_{22}(t)\right)^{-1}\left(\Lambda_{\Sigma} A_{11}(t)+A_{21}(t)\right)  \tag{41}\\
b(t) & =-\frac{\beta}{2} \operatorname{Tr}\left[\log \left(\Lambda_{\Sigma} A_{12}(t)+A_{22}(t)\right)+t\left(M+Q^{\top} \rho z^{\top}\right)\right] \tag{42}
\end{align*}
$$

with

$$
\left(\begin{array}{ll}
A_{11}(t) & A_{12}(t)  \tag{43}\\
A_{21}(t) & A_{22}(t)
\end{array}\right)=\exp t\left(\begin{array}{ll}
M+Q^{\top} \rho z^{\top} & -2 Q^{\top} Q \\
\frac{1}{2}\left(z z^{\top}-\sum_{j=1}^{n} z^{j} e^{j j}\right)+\Lambda_{I} & -\left(M+Q^{\top} \rho z^{\top}\right)^{\top}
\end{array}\right)
$$

In perfect analogy with the WMSV model, also in the WASC model it is possible to compute explicitly the derivative of the function $\alpha \rightarrow G_{\text {WASC }}\left(t, z, \alpha \Lambda_{\Sigma}, \Lambda_{I}\right)$. Since we arrive at the same expression we omit the result.

The remark at the end of section 2.3 on the analytical tractability of the model applies mutatis mutandis here.

## 3 Stock-Volatility Derivative Products

In this section we provide a systematic pricing framework in order to price TVOs, Corridor Variance Swaps and Double Digital Calls within the previously introduced stochastic volatility models. For the TVO our method completes the one proposed e.g. by Torricelli (2013): in fact, it will be clear that a great advantage of our approach is that it is independent of the number of volatility factors. This will be crucial, as we want to apply the methodology to the multi-factor Heston model as well as the Wishart-based stochastic volatility models.

We display separately the results for the different models although the results clearly suggest that we could have unified the presentation. This apparent unity is, however, misleading and we will show later examples for which the Wishart based models or even the BiHeston model introduce strong difficulties.

### 3.1 The Target Volatility Option

The payoff of a Target Volatility Option expiring at time $t$ is given by:

$$
\begin{equation*}
c_{\mathrm{TVO}}=\mathbb{E}\left[\frac{e^{-r t}}{\sqrt{\bar{V}_{t}}} \bar{\sigma}\left(s_{t}-K\right)_{+}\right] \tag{44}
\end{equation*}
$$

with $\bar{V}_{t}=\frac{V_{t}}{t}=\frac{1}{t} \int_{0}^{t} v_{u} d u$ and $\bar{\sigma}$ a positive constant. This contract, in essence, provides the right, but not the obligation, to buy a fractional amount of the stock at the prespecified strike price $K$. The fraction depends on the ratio between the fixed constant $\bar{\sigma}$ and the realized volatility; without loss of generality, we shall set $\bar{\sigma} \equiv 1$. The joint process $\left(s_{t}, v_{t}\right)_{t \geq 0}$ follows e.g. the dynamics given by equations (1) and (2) (we consider for ease of notation the Heston (1993) specification for the volatility process).

First, we express the option price as a function of the Fourier transform of the stock and its volatility. It leads to the following result.

Lemma 3.1. The price of the Target Volatility Option can be expressed as:

$$
c_{\mathrm{TVO}}=\int_{-\infty+i \gamma}^{+\infty+i \gamma} \hat{g}(z) \mathbb{E}\left[\frac{e^{-r t}}{\sqrt{\bar{V}_{t}}} e^{i z \ln s_{t}}\right] d z
$$

where $\hat{g}(z)=-\frac{1}{2 \pi} \frac{K^{1-i z}}{i z(1-i z)}$ and $\gamma=\Im(z)<-1$.
Proof.

$$
\begin{aligned}
\mathbb{E}\left[\frac{e^{-r t}}{\sqrt{\bar{V}_{t}}}\left(s_{t}-K\right)_{+}\right] & =\mathbb{E}\left[\frac{e^{-r t}}{\sqrt{\bar{V}_{t}}} \mathbb{E}\left[\left(s_{t}-K\right)_{+} \mid \mathcal{F}_{v}\right]\right] \\
& =\mathbb{E}\left[\frac{e^{-r t}}{\sqrt{\bar{V}_{t}}} \int_{-\infty}^{+\infty}\left(e^{x}-K\right)_{+} f(x \mid v) d x\right]
\end{aligned}
$$

where $\mathcal{F}_{v}$ is the filtration generated by the volatility path. The density of the logarithm of the stock conditional on the volatility path is given by:

$$
f(x \mid v)=\frac{1}{2 \pi} \int_{-\infty+i \gamma}^{+\infty+i \gamma} e^{-i z x} \mathbb{E}\left[e^{i z \ln s_{t}} \mid \mathcal{F}_{v}\right] d z
$$

Replacing this expression in the previous equation leads, after using Fubini's theorem, to the result. The computation of the Fourier transform $\hat{g}(z)$ is easily done as we have:

$$
\begin{aligned}
\hat{g}(z) & =\frac{1}{2 \pi} \int_{-\infty}^{+\infty} e^{-i z x}\left(e^{x}-K\right)_{+} d x \\
& =\frac{-1}{2 \pi} \frac{K^{1-i z}}{i z(1-i z)}
\end{aligned}
$$

provided that $\Im(z)<-1$, which leads to the constraint on $\gamma$ (see Lewis (2000)).

Remark 3.1. In the previous lemma we consider a call option, which leads to the constraint $\Im(z)<$ -1. Such a constraint can be easily removed by considering a cash-secured put, i.e. a put minus a cash position. Indeed, we have

$$
\begin{equation*}
\frac{1}{2 \pi} \int_{-\infty}^{+\infty} e^{-i z x}\left(\left(K-e^{x}\right)_{+}-K\right) d x=-\frac{1}{2 \pi} \frac{K^{1-i z}}{i z(1-i z)} \tag{45}
\end{equation*}
$$

with $\Im(z) \in]-1,0[$ and the call price in then obtained through call-put parity relation (this remark appears in Lewis (2000) for example). Equivalently, one can considered the generalized Fourier transform of a covered call, which is a call minus a position in the underlying, i.e. $\left(e^{x}-K\right)_{+}-e^{x}$, which gives exactly the same result as in (45). Lemma 3.1 requires the stock to have a moment higher (strictly) than one and is related to the moment explosion problem analyzed in Andersen and Piterbarg (2005).

Hereafter, we will use the well-known relation valid for any $x, \alpha>0$ :

$$
\frac{1}{x^{\alpha}}=\frac{1}{\alpha \Gamma(\alpha)} \int_{0}^{+\infty} e^{-u^{\frac{1}{\alpha}} x} d u
$$

from which we will deduce the price of the TVO for the different models.

Note that this computational trick can be applied with same purpose to certain non affine models, see Leblanc (1996).

TVO in the Heston model For the Heston model we have the following lemma.
Lemma 3.2. In the Heston model of Lemma 2.1 the target volatility option price is:

$$
c_{\mathrm{TVO}}=\frac{2}{\Gamma\left(\frac{1}{2}\right)} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \int_{0}^{+\infty} \hat{g}(z) G_{\mathrm{HES}}\left(t, i z, 0,-\frac{u^{2}}{t}\right) d u d z
$$

where $\hat{g}(z)$ is given in Lemma 3.1 and $\gamma=\Im(z)<-1$.
Proof.

$$
\mathbb{E}\left[\frac{e^{-r t}}{\sqrt{\overline{\bar{V}}_{t}}} e^{i z \ln s_{t}}\right]=\frac{2}{\Gamma\left(\frac{1}{2}\right)} \int_{0}^{+\infty} e^{-r t} \mathbb{E}\left[e^{-u^{2} \bar{V}_{t}+i z \ln s_{t}}\right] d u
$$

The result directly follows by observing that the integrand above depends on the moment generating function $G_{\text {HES }}$ computed in Lemma 2.1.

TVO in the BiHeston model In the BiHeston model we have $\bar{V}_{t}=\frac{V_{t}}{t}=\frac{1}{t} \int_{0}^{t}\left(v_{u}^{0}+v_{u}^{1}\right) d u$, with the following result:

Lemma 3.3. In the BiHeston model of Lemma 2.2 the option price is given by:

$$
c_{\mathrm{TVO}}=\frac{2}{\Gamma\left(\frac{1}{2}\right)} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \int_{0}^{+\infty} \hat{g}(z) G_{2 \mathrm{HES}}\left(t, i z, 0,0,-\frac{u^{2}}{t},-\frac{u^{2}}{t}\right) d u d z
$$

where $G_{2 \mathrm{HES}}$ is the joint Laplace transform defined in Lemma 2.2, $\hat{g}(z)$ is given in Lemma 3.1 and $\gamma=\Im(z)<-1$.

TVO in the WMSV model In the WMSV model we have $\bar{V}_{t}=\frac{V_{t}}{t}=\frac{1}{t} \int_{0}^{t} \operatorname{Tr}\left[\Sigma_{u}\right] d u$ and easily deduce the following result.

Lemma 3.4. In the WMSV model of Lemma 2.3 the option price is given by:

$$
c_{\mathrm{TVO}}=\frac{2}{\Gamma\left(\frac{1}{2}\right)} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \int_{0}^{+\infty} \hat{g}(z) G_{\mathrm{WMSV}}\left(t, i z, 0_{n},-\frac{u^{2}}{t} \mathbb{I}_{n}\right) d u d z
$$

where $G_{\mathrm{WMSv}}$ is the moment generating function defined in Lemma 2.3, $\hat{g}(z)$ is given in Lemma 3.1 and $\gamma=\Im(z)<-1$.

TVO in the WASC model Lastly, for the WASC model, let us consider the payoff of a Target Volatility Option on the first asset $s_{t}^{1}$, such that the option price writes as follows:

$$
c_{\mathrm{TVO}}=\mathbb{E}\left[\frac{e^{-r t}}{\sqrt{\bar{V}_{t}}}\left(s_{t}^{1}-K\right)_{+}\right]
$$

with $\bar{V}_{t}=\frac{V_{t}}{t}=\frac{1}{t} \int_{0}^{t} \Sigma_{u}^{11} d u$. In this case, following the same computations we arrive at the following result.

Lemma 3.5. In the WASC model of Lemma 2.4 the option price is given by:

$$
c_{\mathrm{TVO}}=\frac{2}{\Gamma\left(\frac{1}{2}\right)} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \int_{0}^{+\infty} \hat{g}(z) G_{\mathrm{WASC}}\left(t, i z e^{1}, 0_{n},-\frac{u^{2}}{t} e^{11}\right) d u d z
$$

where $G_{\mathrm{WASC}}$ is defined in Lemma 2.4, $\hat{g}(z)$ is given in Lemma 3.1, $e^{1}$ (resp. $e^{11}$ ) represents the first element of the canonical basis in $\mathbb{R}^{n}$ (resp. in $M_{n}$ ) and $\gamma=\Im(z)<-1$.

### 3.1.1 Remarks regarding the existence of the solutions

We first focus on the Heston model as it is the simplest case and similar arguments will apply to the Bi Heston model. Following Lemma 3.2 we need to check that $\left|G_{\mathrm{HES}}\left(t, i z, 0,-\frac{u^{2}}{t}\right)\right|<+\infty$ where $G_{\mathrm{HES}}$ is given by Lemma 2.1. We have, for $z \in \mathbb{C}$ and $u \in \mathbb{R}:\left|G_{\mathrm{HES}}\left(t, i z, 0,-\frac{u^{2}}{t}\right)\right| \leq G_{\mathrm{HES}}\left(t,-\Im(z), 0,-\frac{u^{2}}{t}\right)$. If we price a call through the call-put parity relation, that is to say we apply Remark 3.1, it leads to consider the function $G_{\text {HES }}\left(t, z, 0,-\frac{u^{2}}{t}\right)$ for $\left.z \in\right] 0,1[$ and $u \in \mathbb{R}$. For these arguments the function $A(t)$ in Lemma 2.1 is given by

$$
A(t)=\frac{\left(z^{2}-z-2 \frac{u^{2}}{t}\right)}{2} \frac{1-e^{-\sqrt{\Gamma} t}}{\lambda_{+}-\lambda_{-} e^{-\sqrt{\Gamma} t}}
$$

with $\Gamma=(\kappa-\rho z \sigma)^{2}-\sigma^{2}\left(z^{2}-z-2 \frac{u^{2}}{t}\right)$ and $\lambda_{ \pm}$as in Equation (3). If $\left.z \in\right] 0,1[$ then $\sqrt{\Gamma}>|\kappa-\rho \sigma|$ and $\lambda_{-}<0$. As we have $\lambda_{+}>0$ then the denominator of $A(t)$ will be strictly positive, thus $A(t)$ is well defined. The function $b(t)$ in the same lemma will also be well defined as it is the integral of $A(t)$. If we wish to apply Lemma 3.2 for a call option (i.e. $\Im(z)<-1$ ) it leads to check that the function $G_{\mathrm{HES}}\left(t, z, 0,-\frac{u^{2}}{t}\right)$ for $z>1$ and $u \in \mathbb{R}$ is well defined. First, for $u$ large enough we have $\Gamma>0$ and $\sqrt{\Gamma}>|\kappa-\rho \sigma z|$ and we can derive the same conclusion as above. If $u=0$ then we are led to consider $\Gamma=(\kappa-\rho z \sigma)^{2}-\sigma^{2}\left(z^{2}-z\right)$. If $z>1$ but still such that $\Gamma>0$ we deduce that $\sqrt{\Gamma}<|\kappa-\rho \sigma z|$. If $\rho<0$ (that will be the case in practice) then $\lambda_{-}>0\left(\right.$ and still $\left.\lambda_{+}>0\right)$ so $t^{*}$ such that $\lambda_{+}-\lambda_{-} e^{-\sqrt{\Gamma} t^{*}}=0$ is equivalent to $t^{*}=-\frac{1}{\sqrt{\Gamma}} \ln \left(\frac{\lambda_{+}}{\lambda_{-}}\right)$. As $\lambda_{+}-\lambda_{-}=\sqrt{\Gamma}>0$ we conclude that $t^{*}<0$, hence there isn't a moment explosion. These results are well known and appear in Andersen and Piterbarg (2005) (to be more precise in this reference the moment of the stock alone is considered
but essentially we follow their procedure). They require an explicit solution for the moment generating function. Although Wishart based stochastic volatility models admit a closed form expression for their moment generating functions, they are not explicit enough so as to allow the analysis performed for the scalar case. However, pricing the call option through the call-put parity relation enables to check the well definiteness of the solution for these multidimensional models. Indeed, for the WMSV model we have $\left|G_{\text {WMSV }}\left(t, i z, 0_{n},-\frac{u^{2}}{t} \mathbb{I}_{n}\right)\right| \leq G_{\mathrm{WMSV}}\left(t,-\Im(z), 0_{n},-\frac{u^{2}}{t} \mathbb{I}_{n}\right)$. Pricing the call through call-put parity relation leads to consider $\Im(z) \in]-1,0\left[\right.$, so to analyze $G_{\mathrm{WMSV}}\left(t, z, 0_{n},-\frac{u^{2}}{t} \mathbb{I}_{n}\right)$ for $\left.z \in\right] 0,1[$. In that case, the matrix Riccati differential equation has the form

$$
\frac{d A}{d t}=A\left(M+z Q^{\top} R^{\top}\right)+\left(M+z Q^{\top} R^{\top}\right)^{\top} A+2 A Q^{\top} Q A+\frac{z(z-1)}{2} \mathbb{I}_{n}-\frac{u^{2}}{t} \mathbb{I}_{n}
$$

with a zero order term that is symmetric negative semi-definite while the quadratic coefficient is symmetric positive definite. Theorem 2.1 of Wonham (1968) ensures that $A(t)$ is well defined and by direct integration we conclude that $b(t)$ given by Equation (28) is also well defined. Regarding the WASC model the same kind of argument can be used as in that case the function $A(t)$ solves, for $z=\left(z^{1}, 0\right)^{\top}$ with $\left.z_{1} \in\right] 0,1[$, the matrix Riccati differential equation

$$
\frac{d A}{d t}=A\left(M+Q^{\top} \rho z^{\top}\right)+\left(M+Q^{\top} \rho z^{\top}\right)^{\top} A+2 A Q^{\top} Q A+\frac{z^{1}\left(z^{1}-1\right)}{2} e^{11}-\frac{u^{2}}{t} e^{11}
$$

and here also the zero order term is symmetric negative semi-definite while the quadratic coefficient is symmetric positive definite.

As a result, it is always possible to perform the integration on the complex plane, through the call-put parity relation, so that the matrix Riccati solutions are well defined. For the Heston model (and BiHeston model) the stock has a moment higher than one. The spot-variance correlation $\rho$ being negative "helps" the process to have higher moments as when the stock goes up its volatility decreases, thus reducing the upper tail of the stock price distribution. For Wishart based models these correlations are also "negative" in the sense that for the WASC both $\rho_{1}$ and $\rho_{2}$ are negative (remember that we have $\rho=\left(\rho_{1}, \rho_{2}\right)^{\top}$ ) leading to a spot-variance correlation that is negative according to Equation (36) while for the WMSV model the spot-variance correlation given by Equation (23) will be negative for market parameter values.

The numerical implementation will be carried out using the call expression given by Lemma 3.1 and for each model the corresponding lemma. Although we are not able to prove that the stock(s) admit a moment higher than one we did not face any explosion problem when pricing call-like options for

Wishart based models (we checked our results using a Monte-Carlo method). We conjecture that this property is true and leave it for future research. In any case, the use of the call-put parity relation is always possible and constitutes an easy-to-implement strategy.

### 3.2 The Corridor Variance Swap

In this sub-section we focus on the pricing of Corridor Variance Swaps, see e.g. Zheng and Kwok (2014), Albanese and Osseiran (2007) and the early work of Carr and Lewis (2004). The payoff at time $t$ is given by

$$
\begin{equation*}
\operatorname{vs}(t)=\mathbb{E}\left[\frac{1}{t} \int_{0}^{t} \mathcal{V}_{u} 1_{\left\{L \leq s_{u} \leq H\right\}} d u-K\right] \tag{46}
\end{equation*}
$$

where $\mathcal{V}_{t}$ is equal to $v_{t}, v_{t}^{0}+v_{t}^{1}, \operatorname{Tr}\left[\Sigma_{t}\right]$ or $\Sigma_{t}^{11}$ depending on which model is considered. The Corridor Variance Swap coincides with a classic Variance Swap provided that the underlying remains in a given corridor defined by the interval $[L, H]$.

The building block for pricing Corridor Variance Swaps is the computation of the term $\mathbb{E}\left[\mathcal{V}_{t} 1_{\left\{x_{t} \leq h\right\}}\right]$ where $x_{t}=\ln s_{t}$ and $h=\ln H$. We have the following result.

Lemma 3.6. Consider $I_{t, h}=\mathbb{E}\left[\mathcal{V}_{t} 1_{\left\{x_{t} \leq h\right\}}\right]$ with $x_{t}=\ln s_{t}$ and $\mathcal{V}_{t}$ defined above. Then we have

$$
\begin{equation*}
I_{t, h}=\frac{1}{2 \pi} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \frac{e^{-i h z}}{-i z} \partial_{\alpha} \mathbb{E}\left[e^{i z x_{t}+\alpha \nu_{t}}\right]_{\left.\right|_{\alpha=0}} d z \tag{47}
\end{equation*}
$$

with $\gamma=\Im(z)>0$.
Proof. We denote by $f\left(x \mid \mathcal{V}_{t}\right)$ the density of $x_{t}$ conditional to $\mathcal{V}_{t}$ and its Fourier transform by $\phi\left(z \mid \mathcal{V}_{t}\right)$ then we have:

$$
\begin{aligned}
I_{t, h} & =\mathbb{E}\left[\mathcal{V}_{t} 1_{\left\{x_{t} \leq h\right\}}\right] \\
& =\mathbb{E}\left[\mathcal{V}_{t} \int_{-\infty}^{+\infty} 1_{\{x \leq h\}} f\left(x \mid \mathcal{V}_{t}\right) d x\right] \\
& =\mathbb{E}\left[\mathcal{V}_{t} \int_{-\infty}^{+\infty} 1_{\{x \leq h\}} \frac{1}{2 \pi} \int_{\mathbb{C}} e^{-i x z} \phi\left(z \mid \mathcal{V}_{t}\right) d z d x\right] \\
& =\frac{1}{2 \pi} \int_{\mathbb{C}} \mathbb{E}\left[\mathcal{V}_{t} \phi\left(z \mid \mathcal{V}_{t}\right)\right] \frac{e^{-i h z}}{-i z} d z \\
& =\frac{1}{2 \pi} \int_{\mathbb{C}} \mathbb{E}\left[\mathcal{V}_{t} e^{i z x_{t}}\right] \frac{e^{-i h z}}{-i z} d z
\end{aligned}
$$

with $\Im(z)>0$. As we have $\mathbb{E}\left[\mathcal{V}_{t} e^{i z x_{t}}\right]=\partial_{\alpha} \mathbb{E}\left[e^{i z x_{t}+\alpha \mathcal{V}_{t}}\right]_{\left.\right|_{\alpha=0}}$ we deduce immediately the result.

Remark 3.2. Similarly, using the above notations we have

$$
\hat{I}_{t, h}=\mathbb{E}\left[\mathcal{V}_{t} 1_{\left\{x_{t} \geq h\right\}}\right]=\frac{1}{2 \pi} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \frac{e^{-i h z}}{i z} \partial_{\alpha} \mathbb{E}\left[e^{i z x_{t}+\alpha \mathcal{V}_{t}}\right]_{\left.\right|_{\alpha=0}} d z
$$

with $\gamma=\Im(z)<0$ that is useful if a different constraint on $\Im(z)$ is needed.
Thanks to the previous lemma and remark we are able to provide the price of the corridor variance swap.

Corollary 3.1. The price of the Corridor Variance Swap is given by

$$
\begin{aligned}
\operatorname{vs}(t) & =\frac{1}{t} \int_{0}^{t} I_{u, h}-I_{u, l} d u-K \\
& =\frac{1}{t} \int_{0}^{t} \hat{I}_{u, l}-\hat{I}_{u, k} d u-K
\end{aligned}
$$

where the quantity $I_{t, h}$ is given in Lemma 3.6, $I_{u, l}$ is similarly defined with $l=\ln L$ while $\hat{I}_{u, h}$ and $\hat{I}_{u, l}$ follow from Remark 3.2.

Finally, the next result expresses the derivative of the moment-generating function according to the different model specification.

Lemma 3.7. The quantity $\partial_{\alpha} \mathbb{E}\left[e^{i z_{1} x_{t}+\alpha \nu_{t}}\right]$ is given by:

$$
\left\{\begin{array}{l}
\partial_{\alpha} G_{\mathrm{HES}}(t, i z, \alpha, 0) \\
\partial_{\alpha} G_{2 \mathrm{HES}}(t, i z, \alpha, \alpha, 0,0) \\
\partial_{\alpha} G_{\mathrm{WMSV}}\left(t, i z, \alpha \mathbb{I}_{n}, 0_{n}\right) \\
\partial_{\alpha} G_{\mathrm{WASC}}\left(t,(i z, 0)^{\top}, \alpha e_{11}, 0_{n}\right)
\end{array}\right.
$$

Let us stress again the fact that the pricing of the corridor variance swap requires the computation of the derivative of the moment-generating function. For all models we consider in this paper (Heston, double Heston, WMSV and WASC) this function is known in closed form. For a standard affine Duffie and Kan (1996) model, for which the Riccati ODEs cannot be explicitly computed (and therefore need to be simulated using a Runge-Kutta scheme for example), it will involve the discretized version of the sensitivity with respect to the initial condition of these Riccati ODEs. In that case the computational burden increases significantly and it underlines the analytical advantages of the Wishart based stochastic volatility models when it comes to build multidimensional extensions. This computational improvement already appears in the pricing of Range Notes; see for example Chiarella et al. (2014) that should be compared with Jang and Yoon (2010). This result also emphasizes the importance of developing alternative expressions for the moment generating function for the WASC and WMSV
models ${ }^{9}$, along these lines see Gnoatto and Grasselli (2014b).

To further illustrate the problem related to the dimension of the state variables let us explain some important differences. In this work we consider the integrated volatility in equation (46) (and follow the definition proposed by Carr and Lewis (2004)) but in practice it is in fact a discretely sampled variance that is traded and its value requires the computation of the quantity:

$$
\mathbb{E}\left[\left(\ln \left(s_{t_{k}}\right)-\ln \left(s_{t_{k-1}}\right)\right)^{2} 1_{\left\{s_{t_{k-1}} \in[L ; H]\right\}}\right]
$$

with $t_{k-1}<t_{k}$. In Zheng and Kwok (2014), to compute such expectation the authors derive twice the moment-generating function with respect to the argument of the stock (its logarithm in fact) and conclude hastily that their "analytic procedure can be applied to any affine model of the underlying asset price and payoff structures of higher moments swaps.". An inspection of the moment-generating functions (24) and (38) shows that their solution applied to these models will lead to more than tedious computations.

### 3.2.1 Remarks regarding the existence of the solutions

The advantage of expressing $\operatorname{vs}(t)$ as a function of $\hat{I}_{u, l}$ and $\hat{I}_{u, k}$ in Corollary 3.1 is the integration constraint required for the computation of these quantities. Focusing first on the Heston model, it involves the function $G_{\text {HES }}(t, i z, 0,0)$ with $\Im(z)<0$. As $\left|G_{\text {HES }}(t, i z, 0,0)\right| \leq G_{\text {HES }}(t,-\Im(z), 0,0)$ it leads to consider the moment generating function $G_{\text {HES }}(t, z, 0,0)$ for $z>0$. Taking $\left.z \in\right] 0,1[$ we know from the previous product that this function is well defined. Similar remark applies to the WMSV and WASC models although well definiteness is checked directly on the matrix Riccati differential equations. As $A(t)$ admits a fraction representation given by Equation (27) the differentiability with respect to the parameter is simple to check. Similar remarks apply to the WASC model. Notice, however, that this simplicity is only due to both the solvability of the Riccati equations of Wishart based models and the very specific dependency of the product on the model parameters (i.e. it involves the derivative with respect to third argument of the moment generating function). Had the product depended on the derivative with respect to the last argument of the moment generating function, that is to say the integrated volatility, the computations would have been far more complicated (see Da Fonseca et al. (2013) for an example of such expressions).

[^7]The numerical implementation will be carried out using the call expression given by Lemma 3.6 and Corollary 3.1. Although we are not able to prove that the stock(s) admit a moment higher than one we did not face any explosion problem for Wishart based models (we checked our results using a Monte-Carlo method). In any case, the use of Remark 3.2 is always possible to keep the integration on the complex plane so that the Riccati's equations are well defined.

### 3.3 The Double Digital Call

In this section we investigate the pricing of a Double Digital Call, see e.g. Torricelli (2013), whose payoff is the indicator function of the event $\left\{s_{t} \geq K_{1}, \bar{V}_{t} \geq K_{2}\right\}$, so that the price is given by

$$
\begin{equation*}
c_{\mathrm{DDC}}\left(s_{0}, K, t\right)=\mathbb{E}\left[e^{-r t} 1_{\left\{s_{t} \geq K_{1}, \bar{V}_{t} \geq K_{2}\right\}}\right] \tag{48}
\end{equation*}
$$

where as usual $\bar{V}_{t}=\frac{V_{t}}{t}=\frac{1}{t} \int_{0}^{t} v_{u} d u$ denotes the integrated variance and $K=\left\{K_{1}, K_{2}\right\}$. We start with the Heston model and check that the results remain valid for multidimensional volatility extensions.

Lemma 3.8. The Double Digital Call option price can be expressed as:

$$
c_{\mathrm{DDC}}=\int_{-\infty+i \gamma}^{+\infty+i \gamma} \hat{g}(z) \mathbb{E}\left[1_{\left\{\bar{V}_{t} \geq K_{2}\right\}} e^{i z \ln s_{t}}\right] d z
$$

where $\hat{g}(z)=\frac{1}{2 i z \pi} e^{-i z \ln K_{1}}$ and $\gamma=\Im(z)<0$.
Proof.

$$
\begin{aligned}
c_{\mathrm{DDC}}\left(s_{0}, K, t\right) & =\mathbb{E}\left[e^{-r t} 1_{\left\{s_{t} \geq K_{1}, \bar{V}_{t} \geq K_{2}\right\}}\right] \\
& =\mathbb{E}\left[e^{-r t} 1_{\bar{V}_{t} \geq K_{2}} \mathbb{E}\left[1_{\left\{s_{t} \geq K_{1}\right\}} \mid \mathcal{F}_{v}\right]\right] \\
& =\mathbb{E}\left[e^{-r t} 1_{\left\{\bar{V}_{t} \geq K_{2}\right\}} \int_{-\infty}^{+\infty} 1_{\left\{e^{x} \geq K_{1}\right\}} f(x \mid v) d x\right]
\end{aligned}
$$

where as usual $\mathcal{F}_{v}$ stands for the filtration generated by the volatility path. The density of the logarithm of the stock conditional on the volatility path is given by:

$$
f(x \mid v)=\frac{1}{2 \pi} \int_{-\infty+i \gamma}^{+\infty+i \gamma} e^{-i z x} \mathbb{E}\left[e^{i z \ln s_{t}} \mid \mathcal{F}_{v}\right] d z
$$

therefore using Fubini's theorem we get:

$$
\begin{aligned}
c_{\mathrm{DDC}}\left(s_{0}, K, t\right) & =\mathbb{E}\left[e^{-r t} 1_{\left\{\bar{V}_{t} \geq K_{2}\right\}} \int_{-\infty}^{+\infty} 1_{\left\{e^{x} \geq K_{1}\right\}} \frac{1}{2 \pi} \int_{\mathbb{C}} e^{-i z x} \mathbb{E}\left[e^{i z \ln s_{t}} \mid \mathcal{F}_{v}\right] d z d x\right] \\
& =e^{-r t} \int_{-\infty}^{+\infty} \hat{g}(z) \mathbb{E}\left[1_{\left\{\bar{V}_{t} \geq K_{2}\right\}} e^{i z \ln s_{t}}\right] d z
\end{aligned}
$$

where

$$
\begin{aligned}
\hat{g}(z) & =\frac{1}{2 \pi} \int_{\mathbb{R}} 1_{\left\{e^{x} \geq K_{1}\right\}} e^{-i z x} d x \\
& =\frac{1}{2 i z \pi} e^{-i z \ln K_{1}}
\end{aligned}
$$

provided that $\Im(z)<0$ in order to grant the convergence of the previous integral.
Now the expression $I\left(s_{0}, v_{0}, t, z\right)=\mathbb{E}\left[1_{\left\{\bar{v}_{t} \geq K_{2}\right\}} e^{i z \ln s_{t}}\right]$ can be computed by using the Laplace-Fourier transform method as presented, among others, in Carr and Madan (1999), Lewis (2000) or Petrella (2004). More precisely, let us consider the Fourier transform of $I\left(s_{0}, v_{0}, t, z\right)$ with respect to $K_{2}$ :

$$
\begin{aligned}
\hat{I}(z, \hat{z}, t) & =\int_{-\infty}^{+\infty} e^{i \hat{z} K_{2}} \mathbb{E}\left[1_{\left\{\bar{V}_{t} \geq K_{2}\right\}} e^{i z \ln s_{t}}\right] d K_{2} \\
& =\mathbb{E}\left[e^{i z \ln s_{t}} \int_{-\infty}^{+\infty} e^{i \hat{z} K_{2}} 1_{\left\{\bar{V}_{t} \geq K_{2}\right\}} d K_{2}\right] \\
& =\mathbb{E}\left[e^{i z \ln s_{t}} \frac{1}{i \hat{z}} e^{i \hat{z} \bar{V}_{t}}\right] \\
& =\frac{1}{i \hat{z}} \mathbb{E}\left[e^{i z \ln s_{t}} e^{\frac{i z}{t} \int_{0}^{t} v_{u} d u}\right] .
\end{aligned}
$$

The expression $\hat{I}(z, \hat{z}, t)$ (with $\Im(\hat{z})<0)$ can be computed explicitly using the joint moment generating function under the different models.

Remark 3.3. We can also rewrite the double digital call as

$$
\begin{align*}
c_{\mathrm{DDC}}\left(s_{0}, K, t\right) & =\mathbb{E}\left[e^{-r t} 1_{\left\{s_{t} \geq K_{1}, \bar{V}_{t} \geq K_{2}\right\}}\right] \\
& =\mathbb{E}\left[e^{-r t} 1_{\left\{s_{t} \geq K_{1}\right\}}\right]-\mathbb{E}\left[e^{-r t} 1_{\left\{s_{t} \geq K_{1}, \bar{V}_{t} \leq K_{2}\right\}}\right] . \tag{49}
\end{align*}
$$

If we denote by $c_{\mathrm{DDC}}^{\mathrm{C}}\left(s_{0}, K, t\right)$ the second expectation in Equation (49) then similar computations as above lead to

$$
c_{\mathrm{DDC}}^{\mathrm{C}}\left(s_{0}, K, t\right)=e^{-r t} \int_{-\infty}^{+\infty} \hat{g}(z) \mathbb{E}\left[1_{\left\{\bar{V}_{t} \leq K_{2}\right\}} e^{i z \ln s_{t}}\right] d z
$$

with $\Im(z)<0$. In that case, $I^{C}\left(s_{0}, v_{0}, t, z\right)=\mathbb{E}\left[1_{\left\{\bar{V}_{t} \leq K_{2}\right\}} e^{i z \ln s_{t}}\right]$ can be computed through LaplaceFourier transform as we have

$$
\hat{I}^{\mathrm{C}}(z, \hat{z}, t)=\frac{1}{i \hat{z}} \mathbb{E}\left[e^{i z \ln s_{t}} e^{\frac{i z}{t}} \int_{0}^{t} v_{u} d u\right] .
$$

with the constraint $\Im(\hat{z})>0$, which is known explicitly. As a consequence, working with $c_{\mathrm{DDC}}^{\mathrm{C}}\left(s_{0}, K, t\right)$ enables different integration paths.

Double Digital Call in the Heston model The expression $\mathbb{E}\left[e^{i z \ln s_{t}} e^{\frac{+i z}{t}} \int_{0}^{t} v_{u} d u\right]$ can be computed explicitly using the joint moment generating function $G_{\text {HES }}\left(t, z, \lambda_{v}, \Lambda_{v}\right)=\mathbb{E}\left[e^{z \ln s_{t}+\lambda_{v} v_{t}+\Lambda_{v} \int_{0}^{t} v_{u} d u}\right]$ defined in Lemma 3, thus giving

$$
\hat{I}(z, \hat{z}, t)=\frac{1}{i \hat{z}} G_{\text {HES }}\left(t, i z, 0, \frac{i \hat{z}}{t}\right) .
$$

Finally, the expression $I\left(s_{0}, v_{0}, t, z\right)$ is given by

$$
\begin{aligned}
I\left(s_{0}, v_{0}, t, z\right) & =\frac{1}{2 \pi} \int_{\mathbb{C}} e^{-i \hat{z} K_{2}} \hat{I}(z, \hat{z}, t) d \hat{z} \\
& =\frac{1}{2 \pi} \int_{\mathbb{C}} e^{-i \hat{z} K_{2}} \frac{1}{i \hat{z}} G\left(t, i z, 0, \frac{i \hat{z}}{t}\right) d \hat{z}
\end{aligned}
$$

and we get the following result.
Lemma 3.9. Under the Heston model of Lemma 2.1 the price of a Double Digital Option is given by:

$$
c_{\mathrm{DDC}}\left(s_{0}, K, t\right)=e^{-r t} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \hat{g}(z) \frac{1}{2 \pi} \int_{-\infty+i \hat{\gamma}}^{+\infty+i \hat{\gamma}} e^{-i \hat{z} K_{2}} \frac{1}{i \hat{z}} G_{\mathrm{HES}}\left(t, i z, 0, \frac{i \hat{z}}{t}\right) d \hat{z} d z,
$$

with $\gamma=\Im(z)<0, \hat{\gamma}=\Im(\hat{z})<0$.
Double Digital Call in the BiHeston model The expression $\mathbb{E}\left[e^{i z \ln s t} e^{\frac{+i z}{t} \int_{0}^{t}\left(v_{u}^{0}+v_{u}^{1}\right) d u}\right]$ can be computed explicitly using the moment generating function $G_{2 \text { HES }}$ defined in Lemma 2.2, thus giving

$$
\hat{I}(z, \hat{z}, t)=\frac{1}{i \hat{z}} G_{2 \mathrm{HES}}\left(t, i z, 0,0, \frac{i \hat{z}}{t}, \frac{i \hat{z}}{t}\right) .
$$

Lemma 3.10. Under the BiHeston model of Lemma 2.2 the price of a Double Digital Option is given by:

$$
c_{\mathrm{DDC}}\left(s_{0}, K, t\right)=e^{-r t} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \hat{g}(z) \frac{1}{2 \pi} \int_{-\infty+i \hat{\gamma}}^{+\infty+i \hat{\gamma}} e^{-i \hat{z} K_{2}} \frac{1}{i \hat{z}} G_{2 \mathrm{HES}}\left(t, i z, 0,0, \frac{i \hat{z}}{t}, \frac{i \hat{z}}{t}\right) d \hat{z} d z,
$$

with $\gamma=\Im(z)<0, \hat{\gamma}=\Im(\hat{z})<0$.
Double Digital Call in the WMSV model The expression $\mathbb{E}\left[e^{i z \ln s_{t}} e^{\frac{i z}{t}} \int_{0}^{t} \operatorname{Tr}\left[\Sigma_{u}\right] d u\right]$ can be computed explicitly using the moment generating function $G_{\mathrm{WMSv}}\left(t, z, \Lambda_{\Sigma}, \Lambda_{I}\right)$ defined in Lemma 2.3 , thus giving the following result.

Lemma 3.11. Under the WMSV model Lemma 2.3 the price of a Double Digital Option is given by:

$$
c_{\mathrm{DDC}}\left(s_{0}, K, t\right)=e^{-r t} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \hat{g}(z) \frac{1}{2 \pi} \int_{-\infty+i \hat{\gamma}}^{+\infty+i \hat{\gamma}} e^{-i \hat{z} K_{2}} \frac{1}{i \hat{z}} G_{\mathrm{WMSV}}\left(t, i z, 0_{n}, \frac{i \hat{z}}{t} \mathbb{I}_{n}\right) d \hat{z} d z,
$$

with $\gamma=\Im(z)<0, \hat{\gamma}=\Im(\hat{z})<0$.

Double Digital Call in the WASC model The expression $\mathbb{E}\left[e^{i z \ln s_{t}^{1}} e^{\frac{i z}{t}} \int_{0}^{t} \sum_{u}^{11} d u\right]$ can be computed explicitly using the moment generating function $G_{\text {wasc }}\left(t, z, \Lambda_{\Sigma}, \Lambda_{I}\right)$ defined in Lemma 2.4, thus giving the following result.

Lemma 3.12. Under the WASC model of Lemma 2.4 the price of a Double Digital Option on the first asset is given by:

$$
c_{\mathrm{DDC}}\left(s_{0}^{1}, K, t\right)=e^{-r t} \int_{-\infty+i \gamma}^{+\infty+i \gamma} \hat{g}(z) \frac{1}{2 \pi} \int_{-\infty+i \hat{\gamma}}^{+\infty+i \hat{\gamma}} e^{-i \hat{z} K_{2}} \frac{1}{i \hat{z}} G_{\mathrm{WASC}}\left(t, i z e_{1}, 0_{n}, \frac{i \hat{z}}{t} e_{11}\right) d \hat{z} d z,
$$

with $\gamma=\Im(z)<0, \hat{\gamma}=\Im(\hat{z})<0$.

### 3.3.1 Remarks regarding the existence of the solutions

First, we consider the Heston model as the same remarks will apply to the BiHeston model. As we $\left|G_{\text {HES }}\left(t, i z, 0, \frac{i \hat{z}}{t}\right)\right| \leq G_{\text {HES }}\left(t,-\Im(z), 0,-\frac{\Im(\hat{z})}{t}\right)$, if we price the double digital call using Remark 3.3 it leads to consider $G_{\text {HES }}\left(t, z, 0,-\frac{\hat{z}}{t}\right)$ for $z>0$ and $\hat{z}>0$. The term $\Gamma$ of Equation (4) will have the expression $\Gamma=(\kappa-\rho \sigma z)^{2}-\sigma^{2}\left(z(z-1)-\frac{2 \hat{z}}{t}\right)$. Following previous discussion if $\left.z \in\right] 0,1[$ and $\hat{z}>0$ then we conclude that there is no explosion. If we consider the WMSV model, the zero order term of the matrix differential Riccati equation is $\frac{z(z-1)}{2} \mathbb{I}_{n}-\frac{\hat{z}}{t} \mathbb{I}_{n}$, thus it will be a symmetric negative definite matrix while for the WASC model it will be $\frac{z^{1}\left(z^{1}-1\right)}{2} e^{11}-\frac{\hat{z}}{t} e^{11}$ that is symmetric negative semi-definite matrix. In both cases, Theorem 2.1 in Wonham (1968) ensures that there is no explosion.

## 4 Numerical Results

We implement the formulas presented in the previous parts and provide for each products some practical details. The pricing of the target volatility option involves a two dimensional integration, for the Heston model it is given by Lemma 3.2 but similar remarks apply to the other models, that will be discretized using Simpson's rule scheme. The inner integral, whose integration variable is $u$, will be restricted to the interval $[0,20]$ partitioned with 256 regularly spaced points while for the outer integral the integration will be carried out on $[-5,5]$ partitioned with 256 regularly spaced points. Furthermore, we will take $\gamma=\Im(z)=-1.2$, thus smaller than -1 as required. For the corridor variance swap only a one-dimensional integration is needed. As the left hand side of Equation (47) is real we rewrite this integral as

$$
I_{t, h}=\frac{1}{\pi} \int_{0+i \gamma}^{+\infty+i \gamma} \Re\left(\frac{e^{-i h z}}{-i z} \partial_{\alpha} \mathbb{E}\left[e^{i z x_{t}+\alpha \mathcal{V}_{t}}\right]_{\left.\right|_{\alpha=0}}\right) d z
$$

We discretize the above integral using a simple trapezoid rule with 4096 points regularly spaced by $\delta=0.18$ and with the first point being at 0 . We take $\gamma=\Im(z)=0.3$ to be consistent with Lemma 3.6. The above procedure produces an array which stores evaluations of the integrand under study, which could be then processed via and FFT routine, thus allowing for the computation of $I_{t, h}$ for several values of $h^{10}$. Lastly, for the double digital call option the two-dimensional Fourier transform of Lemma 3.9 is computed by discretization of the integrals. The inner integral will be restricted to $\left[-\frac{\bar{Y}}{2}, \frac{\bar{Y}}{2}\right]$ with $\bar{Y}=N d \omega, d \omega=0.1$ and $N=40000$ while for the other integral it will be $\left[-\frac{\bar{X}}{2}, \frac{\bar{X}}{2}\right]$ with $\bar{X}=N d x, d x=0.0025$. We take $\gamma=-1$ and $\hat{\gamma}=-0.01$ both negative as required by lemmas 3.9 , $3.10,3.11$ and 3.12 . The shape of the integrand allows the reduction of the number of characteristic function evaluations. Being integrable, the integrand converges to 0 when the argument's modulus converges to infinity. In our case, it leads to start the computation of the integral from the center of the grid (i.e. close to 0 ) and move away from it, the integration along a line is stopped whenever the modulus of the integrand is small enough. For example, in Lemma 3.9, which applies to the Heston model, the condition will be

$$
\left|\frac{1}{2 \pi} \hat{g}(z) \frac{e^{-i z K_{2}}}{i \hat{z}} G_{\mathrm{HES}}\left(t, i z, 0, \frac{i \hat{z}}{t}\right) d \omega d x\right|<10^{-15}
$$

Similar rule is applied to the other models. This computational property reveals to be highly relevant for fast pricing in the Wishart-based models.

The four models were calibrated on same data so they produce approximately the same vanilla option values, they are extracted from Da Fonseca and Grasselli (2011) and reported in Table I. Naturally, models with more parameters lead to a smaller calibration error but even for the Heston model the pricing error is relatively small. For the values presented here the root mean square error for out-themoney option prices is $0.163 \%$ of the underlying forward price while for the BiHeston, WMSV and WASC models it is around $0.1 \%$. The pricing of these exotic options with calibrated models allows us to put our results in the broader perspective of model risk, see Cont (2006) for related aspects, and raises some practical important problems.
[ Insert Table I here]

The TVO prices are reported in Table II, the Corridor Variance Swap prices are reported in Table III while the double digital call prices are given, depending on the model considered, by Tables IV, V, VI and VII.

[^8]For the TVO prices the Heston and BiHeston models lead to similar prices. However, note that the percentage difference between the prices given by the two models can reach $10 \%$ (e.g., for the maturity 0.5 and strike 0.9 ) and on average around $5 \%$. For the WASC and WMSV models (for both models we changed the Gindikin parameters so that they are greater than one) the average discrepancy between the TVO prices is $3 \%$ and decreases with the maturity.

For the Corridor Variance Swap the Heston and BiHeston models give prices that are close but the error increases with the maturity ( $0.8 \%$ for the maturity 0.5 and $5 \%$ for the one year maturity). For the WASC and WMSV, we observe the opposite, that is the average discrepancy decreases with the maturity, from $10 \%$ to $6 \%$.

For the Double Digital Call the conclusions are similar. For example, comparing the Heston and BiHeston models illustrates the fact that models giving close vanilla prices can lead to substantial differences in derivative prices as a difference of more than $50 \%$ can easily be reached if we consider options whose payoff depends on the tail of the asset-volatility distribution. This problem is likely to be magnified by complex payoff structures. Similarly, the differences between the WASC and the WMSV can be substantial ( $30 \%$ for $T=1, K_{1}=0.08$ and $K_{1}=1$ ). Also of interest is the fact that for the WASC and WMSV we changed slightly the Gindikin parameters, which gives us a rough idea of exotic option prices sensitivity to model parameter "uncertainty" and illustrates how prices produced by the pair Heston/BiHeston and the pair WMSV/WASC can diverge for small parameter perturbations. Because the DDC strongly depend on the tails of the stock-volatility distribution it constitutes a "worst" case example, the other products lead to similar, though less dramatic conclusions. A comparison between the Heston/BiHeston prices on one hand and the WMSV/WASC prices on the other hand shows a huge difference. Let us stress the fact that the magnitude of our values are in line with what happens in practice. It should be clear also that adding exotic options in the calibration objective function, so that all the models produce similar vanilla and exotic option prices, does not guarantee that exotic options not included in the calibration set will be similarly priced. Furthermore, this strategy supposes that these exotic option prices are given, that is to say are input values and implies that these products are liquid enough, which is often not the case.

As noted in Chiarella et al. (2014), in practice whenever a derivatives seller wants to gain some
confidence in his pricing the standard procedure is to ask other market participants for their price. As some derivatives can be very exotic it might be difficult to obtain such information. To overcome this difficulty some companies provide a service that allows a market participant to know whether his price is close or within the range of prices proposed by the other participants (but without revealing the prices). Our equity-volatility option results confirm the issues raised with this practice and extend to this market the concerns developed in Chiarella et al. (2014) for the interest rate markets. It is unclear to us whether current market regulation rules address properly that problem.

## 5 Remarks and Open Problems

The previous results might suggest that any closed-form results for the Heston model can be easily extended to the BiHeston, WMSV or WASC models. We already mentioned that this statement is not correct. In addition to the problems underlined in the corridor variance swap section, the pricing of option on the discretely sampled variance once more illustrates the difficulty the handle multivariate models. It was performed in the Heston model in Lian et al. (2014) and it is simple to check that adapting to the WMSV or WASC models their results is a non-trivial task. Let us further illustrate with another equity-stochastic volatility product the difficulties related to the dimension. The timer option is a recent product whose payoff is given by

$$
\mathbb{E}\left[e^{-r_{\tau_{t}}}\left(s_{\tau_{t}}-K\right)_{+}\right]
$$

with $\tau_{t}=\inf \left\{u ; \int_{0}^{u} v_{s} d s=t\right\}$ and $v_{u}$ is the volatility of the stock. As for the products considered in this work the timer option payoff depends on the stock and its volatility (path). For this product a closed-form solution is available for the Heston model, see Li (2013), but the extension to the WMSV or WASC model leads to important difficulties that we were not able to solve. Even the BiHeston, which does not involve any matrices in its characteristic function, brings some tedious numerical difficulties.

Lastly, we only consider continuous time diffusion processes while the univariate stochastic volatility model can have different kind of jumps (on the stock and/or the volatility). An alternative to the Wishart-based models was proposed in the series of papers Barndorff-Nielsen and Stelzer (2013) and Muhle-Karbe et al. (2012) who develop a multi-asset matrix jump process model, the pricing of exotic derivatives within that framework is an open question ${ }^{11}$.

[^9]
## 6 Conclusion

In this paper we exploited a powerful technique in order to price some equity-volatility products that have been recently introduced in the market. Our approach combines conditioning with respect to the subfiltration generated by the volatility path with simple Fourier techniques. Our methodology allows one to price in closed form Target Volatility options, Corridor Variance Swaps and Double Digital calls regardless of the dimension of the stochastic process used to describe the volatility process. We investigated the affine class with a special emphasis on the recent Wishart based specifications introduced by Da Fonseca, Grasselli, and Tebaldi (2007) and Da Fonseca, Grasselli, and Tebaldi (2008), for which closed form solutions are available for the moment generating function.

A numerical exercise for the TVO and Corridor Variance Swap shows that the Heston and BiHeston models lead to similar prices for short maturities; however the discrepancy between the prices increases with the maturity and can reach $10 \%$ but is on average around $5 \%$. For the WASC and WMSV models we observe the opposite, that is the average discrepancy between the prices is $3 \%$ and decreases with the maturity. These results suggest that models giving close vanilla prices can lead to substantial differences in exotic derivative prices. Furthermore, a small perturbation of data parameters can produce huge differences for exotic derivative prices and raises the question of how to define a robust pricing for these derivatives.

Our results clearly illustrate the remarkable flexibility of Wishart-based models as they enable to increase the dimension of the state variables, either of the volatility or the number of assets, and yet remain highly tractable. Although we showed how to overcome some difficulties we pointed out some open and challenging questions that we leave for future research.

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## Appendix

## Tables

Table I: Model Parameter Values

| Heston | Value | BiHeston | Value | WASC | Value | WSMV | Value |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $v_{t}$ | 0.0414 | $v_{t}^{0}$ | 0.0187 | $\Sigma_{t}^{11}$ | 0.0446 | $\Sigma_{t}^{11}$ | 0.0298 |
| $\kappa$ | 1.4078 | $\kappa_{0}$ | 1.3080 | $\Sigma_{t}^{12}$ | 0.0366 | $\Sigma_{t}^{12}$ | 0.0119 |
| $\theta$ | 0.0838 | $\theta_{0}$ | 0.0281 | $\Sigma_{t}^{22}$ | 0.0424 | $\Sigma_{t}^{22}$ | 0.0108 |
| $\sigma$ | 0.9319 | $\sigma_{1}$ | 1.1202 | $\beta$ | 1.7332 | $\beta$ | 1.5776 |
| $\rho$ | -0.5409 | $\rho_{0}$ | -0.3884 | $M_{11}$ | -0.7820 | $M_{11}$ | -1.2479 |
|  |  | $v_{t}^{1}$ | 0.0229 | $M_{12}$ | -0.3772 | $M_{12}$ | -0.8985 |
|  |  | $\kappa_{1}$ | 1.4134 | $M_{21}$ | -0.0539 | $M_{21}$ | -0.0820 |
|  |  | $\theta_{1}$ | 0.0485 | $M_{22}$ | -1.2497 | $M_{22}$ | -1.1433 |
|  |  | $\sigma_{1}$ | 0.4822 | $Q_{11}$ | 0.3898 | $Q_{11}$ | 0.3417 |
|  |  | $\rho_{1}$ | -0.8395 | $Q_{12}$ | 0.3573 | $Q_{12}$ | 0.3493 |
|  |  |  |  | $Q_{21}$ | 0.2809 | $Q_{21}$ | 0.1848 |
|  |  |  |  | $Q_{22}$ | 0.3362 | $Q_{22}$ | 0.3090 |
|  |  |  |  | $\rho_{1}$ | -0.6407 | $R_{11}$ | -0.2243 |
|  |  |  |  | $\rho_{2}$ | -0.1105 | $R_{12}$ | -0.1244 |
|  |  |  |  |  |  | $R_{21}$ | -0.2545 |
|  |  |  |  |  |  | $R_{22}$ | -0.7230 |

These parameter values are those of Da Fonseca and Grasselli (2011) and were obtained by performing a calibration on the DAX vanilla options on the day August, 20 2008. For the WASC model the calibration was performed on the options for the pair EuroStoxx50/DAX. Without loss of generality we will take $s_{0}=1$ and the risk free rate $r=0$. We changed the Gindikin parameter values $\beta$ so that $\beta>1$.

Table II: TVO Prices

|  | Heston |  |  | BiHeston |  | WASC |  | WMSV |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Maturity | 0.5 | 1.0 | 0.5 | 1.0 | 0.5 | 1.0 | 0.5 | 1.0 |  |
|  | 0.9 | 0.93084 | 0.98910 | 0.84322 | 0.92055 | 0.58416 | 0.66302 | 0.62240 | 0.68281 |
| K | 1 | 0.56497 | 0.63592 | 0.51799 | 0.60051 | 0.38502 | 0.48436 | 0.40299 | 0.48864 |
|  | 1.1 | 0.30230 | 0.37550 | 0.28313 | 0.36242 | 0.23660 | 0.34307 | 0.24106 | 0.33748 |

We report the TVO price $c_{\text {Tvo }}$ for the maturity $t \in\{0.5,1\}$, the strike $K \in\{0.9,1,1.1\}$ and model parameter values given in Table I.

Table III: Corridor Variance Swap Prices

|  | Heston |  | BiHeston |  | WASC |  | WMSV |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Maturity |  | 0.25 | 0.75 | 0.25 | 0.75 | 0.25 | 0.75 | 0.25 | 0.75 |
| K | $[0.95$ | $1.05]$ | 0.014354 | 0.010273 | 0.014462 | 0.009730 | 0.025362 | 0.018026 | 0.028327 |
|  | $[0.9$ | $1.1]$ | 0.025931 | 0.019664 | 0.026170 | 0.018824 | 0.047200 | 0.035517 | 0.053169 |
|  |  |  |  | 0.038036 |  |  |  |  |  |

We report the quantities $I_{t, h}-I_{t, l}$ for $t \in\{0.5,1\},[L, H]$ equals to [0.95 1.05] or [0.9 1.1] for the four models and parameter values given in Table I.

Table IV: Double Digital Call Prices - Heston

| $T=0.5$ |  |  |  |  |  | $T=1$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $K_{1}$ | 0.9 | 1 | 1.1 | 0.9 | 1 | 1.1 |  |
| $K_{2}$ | 0.03 | 0.19901 | 0.13377 | 0.06070 | 0.18938 | 0.13806 | 0.08370 |  |
|  | 0.05 | 0.10889 | 0.06956 | 0.03737 | 0.10493 | 0.07440 | 0.04845 |  |
|  | 0.08 | 0.03186 | 0.01788 | 0.01475 | 0.03066 | 0.02049 | 0.01655 |  |

We report the double digital call prices for the maturity $T \in\{0.5,1\}, K_{1} \in\{0.9,1,1.1\}$ and $K_{2} \in\{0.03,0.05,0.08\}$ for the Heston model with parameter values given in Table I.

Table V: Double Digital Call Prices - BiHeston

|  | $T=0.5$ |  |  |  | $T=1$ |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $K_{1}$ | 0.9 | 1 | 1.1 | 0.9 | 1 | 1.1 |
|  | 0.03 | 0.21303 | 0.13979 | 0.05990 | 0.20186 | 0.14483 | 0.08550 |
| $K_{2}$ | 0.05 | 0.11407 | 0.06842 | 0.03136 | 0.10982 | 0.07388 | 0.04326 |
|  | 0.08 | 0.02044 | 0.00755 | 0.00688 | 0.02049 | 0.01001 | 0.00673 |

We report the double digital call prices for the maturity $T \in\{0.5,1\}, K_{1} \in\{0.9,1,1.1\}$ and $K_{2} \in\{0.03,0.05,0.08\}$ for the BiHeston model with parameter values given in Table I.

Table VI: Double Digital Call Prices - WMSV

|  |  | $T=0.5$ |  |  |  | $T=1$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $K_{1}$ | 0.9 | 1 | 1.1 | 0.9 | 1 | 1.1 |
|  | 0.03 | 0.32062 | 0.23711 | 0.14847 | 0.29719 | 0.23862 | 0.18013 |
| $K_{2}$ | 0.05 | 0.26166 | 0.18894 | 0.11706 | 0.26909 | 0.21324 | 0.15919 |
|  | 0.08 | 0.15828 | 0.10818 | 0.06565 | 0.20056 | 0.15521 | 0.11421 |

We report the double digital call prices for the maturity $T \in\{0.5,1\}, K_{1} \in\{0.9,1,1.1\}$ and $K_{2} \in\{0.03,0.05,0.08\}$ for the WMSV model with parameter values given in Table I.

Table VII: Double Digital Call Prices - WASC

|  | $T=0.5$ |  |  |  |  | $T=1$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $K_{1}$ | 0.9 | 1 | 1.1 | 0.9 | 1 | 1.1 |
|  | 0.03 | 0.29818 | 0.21497 | 0.12874 | 0.29105 | 0.22963 | 0.16940 |
| $K_{2}$ | 0.05 | 0.22936 | 0.16139 | 0.09694 | 0.24292 | 0.18912 | 0.13861 |
|  | 0.08 | 0.12577 | 0.08412 | 0.05112 | 0.15466 | 0.11719 | 0.08500 |

We report the double digital call prices for the maturity $T \in\{0.5,1\}, K_{1} \in\{0.9,1,1.1\}$ and $K_{2} \in\{0.03,0.05,0.08\}$ for the WASC model with parameter values given in Table I.


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[^1]:    ${ }^{1}$ Let us just mention, without pretending to be exhaustive, some works from this growing literature; Sepp (2008), Bao et al. (2012), Zhu and Zhang (2007), Shen and Siu (2013) and Lian et al. (2014).
    ${ }^{2}$ To be more precise some authors consider the Heston model with jumps on the stock and/or the volatility, with the unifying computational framework proposed in Duffie et al. (2000), but for the products considered here the jumps do not introduce any special difficulty. We refer to these jump-diffusion models as the Heston model although in its initial specification it has no jumps

[^2]:    ${ }^{3}$ Hereafter we consider the unconditional moment generating function and we will provide the price at time zero of a payoff maturing at time $t$. Of course in our homogeneous Markovian setting we can easily adapt the arguments to the conditional moment generating function $G_{\text {HES }}\left(t, z, \lambda_{v}, \Lambda_{v}\right)=\mathbb{E}_{s}\left[e^{z \ln s_{t}+\lambda_{v} v_{t}+\Lambda_{v} \int_{0}^{t} v_{u} d u}\right]$, for $s \leq t$.

[^3]:    ${ }^{4}$ In other words, $d W_{t}^{0} d W_{t}^{1}=d Z_{t}^{0} d Z_{t}^{1}=d W_{t}^{0} d Z_{t}^{1}=d W_{t}^{1} d Z_{t}^{0}=0$ in order to grant the affine property of the infinitesimal generator, see e.g. Da Fonseca et al. (2008).

[^4]:    ${ }^{5}$ For other option pricing applications of this model see for example Benabid et al. (2008), Branger and Muck (2012), Leung et al. (2013) and Gnoatto and Grasselli (2014a).
    ${ }^{6}$ The constraint on $\Omega \Omega^{\top}$ can be relaxed by requiring that $\Omega \Omega^{\top}-(n-1) Q^{\top} Q$ is positive semidefinite, as explained in Cuchiero et al. (2011). We keep our more parsimonious choice in view of numerical applications as to the best of our knowledge it is the only specification for which a calibration on market option prices is available.

[^5]:    ${ }^{7}$ To simplify notations we omit the dependency of these functions on the time variable $t$ in the ODE.

[^6]:    ${ }^{8}$ If we were to define a multiple-asset model using several simple Heston models then introducing some correlation between the assets will turn the model non affine.

[^7]:    ${ }^{9}$ The formulas for these two models are obtained through linearization of Riccati's equations, as suggested by Grasselli and Tebaldi (2008).

[^8]:    ${ }^{10}$ The standard radix-2 algorithm is used through the GSL library (GNU Scientific Library).

[^9]:    ${ }^{11}$ There are only very few multiasset stochastic volatility models, apart from those mentioned above that have the feature of employing matrix diffusion processes let us also mention Yoon et al. (2011).

